

## THE RELATIONSHIP BETWEEN INVESTMENT AND GDP: AN ARDL APPROACH FOR THE CASE OF IRAN

Mohsen Mehrara\*

### *Abstract*

The objective of this paper is to examine the relationship between investment and economic growth in Iran for the period 1970-2010, based on the autoregressive distributed lag (ARDL) approach. The study finds a cointegrating relationship among real GDP, investment, labor force, oil revenues and education. Compared to the other variables, labor force and human capital (education) have the most important effect on long-run economic growth. Moreover, in short-run oil revenues and investment have the strongest effects on economic growth. Estimating error correction model revealed that the speed of adjustment to restore equilibrium is -0.71 which confirms that there is a stable long-run relationship. Regarding weak impact of investment on long run economic growth, it seems that government and policy makers should employ policies that would accelerate economic growth through higher productivity and privileged human capital.

**JEL classification:** C12, C22, C52, E21, F43

**Keywords:** ARDL, Gross Domestic Investment, Economic Growth, Iran Economy

---

\* Faculty of Economics, University of Tehran, Tehran, Iran

### *1. Introduction*

In the theories of economic growth, it is often argued that increased level of investment can enhance economic growth. According to theories of investment, on the other hand, investment can be affected by many factors, including economic growth. Hence, if economic growth rises, the investment increase too. According to these two point of views expressed in economic theories, investigating the relationship between investment and economic growth is an important as well as controversial issue for economists and policy makers. Some researchers have analyzed it as cause and effect relationship. Another group of economists believe in capital fundamentalists' point of view that investment cause growth. There is a third group of economics who have confirmed the theory that investment depends upon the level of output.

Summers (1991, 1992), studying the relationship between investment to GDP ratio and growth rates since World War II; conclude that the investment determines the rate of economic growth. Solow (1956) in his study indicates that the larger investment and saving rate lead to more output per worker. Tyler (1981) investigating a sample of 55 developing countries showed that investment is the main determinant of growth rate. New growth theories insist on the importance of investments, human and physical capital in the long-run economic growth. The policies, which affect the level of growth and the investment productivity, determine the long-run economic growth. Theoretically, the gross investment affects the economic growth through either increasing the physical capital stock, or promoting the technology (Levine and Renelt, 1992 and Plossner, 1992). However, Blomstrom et al. (1996), based on econometric results of the Granger causality tests indicate that the direction of causality runs in the opposite way. That is, it is the GDP growth that leads to more investment in the economy. Therefore, they conclude something else except investment should explain economic growth including economic and political environment, productivity, education, foreign investment and so on.

As a result, because of the importance of investment in economic growth and also the effective role of economic growth in increasing investment, it is necessary to investigate the relationship between the level of investment and economic growth in Iran, and its results can be useful in designing policies.

In this paper we examine the short- and long-run relationships between investment and economic growth for Iran over the period 1970-2010, using Autoregressive Distributed Lag (ARDL) approach to cointegration and error correction models (ECM).

The rest of the paper is organized as follows. Section 2 describes data and methodology. Results are reported in Section 3. Section 4 concludes.

## 2. Data and Methodology

To allow for causality and dynamics and given that not all of our time-series may be stationary to the same order (some are I(0) while others are I(1)), the cointegration technique suggested by Pesaran et al. (2001), the autoregressive distributed lag model (ARDL) procedure will be used. The approach can be implemented regardless of whether the variables are integrated of order (1) or (0) and can be applied to small finite samples. Based on empirical literature, theories of economic growth, and diagnostic tests, the long run relationship between economic growth and investment can be specified as:

$$\ln RGDP_t = \beta_0 + \beta_1 \ln INV_t + \beta_2 \ln L_t + \beta_3 \ln OILREV_t + \beta_4 SER_t + u_t \quad (1)$$

Where RGDP is GDP at constant price, INV is gross domestic investment, L is labor force, OILREV is real oil revenues, SER is the secondary enrolment ratio and proxies for the quality of human capital.  $\varepsilon_t$  is an stationary error term. All variables except SER are expressed in natural logarithm (ln stands for logarithm). The main sources of variables are from the Central Bank of Iran (CBI) and Statistical Center of Iran (SCI). The time period of the study is over the years 1970 to 2010.

To examine long run relation among the series we implement ARDL bounds testing approach to cointegration developed by Pesaran et al., (2001). The bounds testing approach has several advantages: it applies irrespective of the order of integration for independent variables, I(0) or I(1); is better suited to small samples; and a dynamic error correction model (ECM) can be derived from the ARDL model through a simple linear reparametrization. The version of error correction model of ARDL approach is given by:

$$\begin{aligned} \Delta \ln RGDP_t = & \alpha_0 + \sum_{i=1}^p \phi_i \Delta \ln RGDP_{t-i} + \sum_{i=0}^p \theta_i \Delta \ln INV_{t-i} + \sum_{i=0}^p \lambda_i \Delta \ln L_{t-i} \\ & + \sum_{i=0}^p \varphi_i \Delta \ln OILREV_{t-i} + \sum_{i=1}^p \gamma_i \Delta \ln SER_{t-i} + \delta_1 \ln RGDP_{t-1} + \delta_2 \ln INV_{t-1} \\ & + \delta_3 \ln L_{t-1} + \delta_4 \ln OILREV_{t-1} + \delta_5 SER_{t-1} + \varepsilon_t \end{aligned} \quad (2)$$

Where  $\phi, \theta, \lambda, \rho$  and  $\gamma$  refer to short run and  $\delta_1$  to  $\delta_5$  to long run parameters. The null hypothesis of no cointegration is  $H_0: \delta_1 = \delta_2 = \delta_3 = \delta_4 = \delta_5 = 0$  against the alternative hypothesis  $H_1: \delta_1 \neq \delta_2 \neq \delta_3 \neq \delta_4 \neq \delta_5 \neq 0$ . The rejection of the null based on the F-statistic suggests cointegrating relationship. The critical bounds have been tabulated by Pesaran et al. (2001). The upper critical bound (UCB) is based on the assumption that all series are I(1). The lower bounds (LCB) applies if the series are I(0). If UCB is lower than the calculated F-statistic, the null of cointegration is sustained. If the F-statistic is less than the LCB then there is no cointegration. The decision about cointegration will be inconclusive if the F-statistic lies between UCB and LCB. In such situation, we will have to rely on the lagged error correction term to investigate long run relationship.

The orders of the lags in the specification (2) are selected by the Schwarz Bayesian criterion (SBC). For annual data, Pesaran and Shin (1999) recommended choosing a maximum of 2 lags. From this, the lag length that minimizes SBC is selected.

If a long run relationship exists, the ARDL representation of equation (1) is formulated as follows:

$$\ln RGDP_t = \alpha_1 + \sum_{i=1}^{p+1} \phi_{1i} \ln RGDP_{t-i} + \sum_{i=1}^{p+1} \rho_{1i} \ln INV_{t-i} + \sum_{i=1}^{p+1} \theta_{1i} \ln L_{t-i} + \sum_{i=0}^{p+1} \lambda_{1i} \ln OILREV_{t-i} + \sum_{i=1}^{p+1} \varphi_{1i} SER_{t-i} + \varepsilon_t \quad (3)$$

The ARDL method estimate  $(p+1)^k$ , number of regressions in order to obtain the optimal lags for each variable, where  $p+1$  is the maximum number of lags to be used and  $k$  is the number of variables in the equation (Shrestha and Chowdhury, 2005). The model is selected based on the Schwartz-Bayesian Criterion (SBC) that use the smallest possible lag length and is therefore described as the parsimonious model.

The ARDL specification of short run dynamics is investigated using ECM version of ARDL model of the following form:

$$\Delta \ln RGDP_t = \alpha_2 + \sum_{i=1}^p \phi_{2i} \Delta \ln RGDP_{t-i} + \sum_{i=1}^p \rho_{2i} \Delta \ln INV_{t-i} + \sum_{i=1}^p \theta_{2i} \Delta \ln L_{t-i} + \sum_{i=1}^p \lambda_{2i} \Delta \ln OILREV_{t-i} + \sum_{i=0}^p \varphi_{2i} \Delta SER_{t-i} + \psi ECM_{t-1} + \varepsilon_t \quad (4)$$

The lagged residual term (ECM) in equation 4 shows the disequilibrium in long run relationship (ut in equation 1). The goodness of fit for ARDL model is checked through stability tests such as cumulative sum of recursive residuals (CUSUM) and cumulative sum of squares of recursive residuals (CUSUMSQ).

### 3. Empirical Results

Pesaran et al. (2001) critical values are based on the assumption that the variables are integrated of order I(0) or I(1). Unit root tests insure that none of the series is integrated of I(2) or higher. Both the augmented Dickey–Fuller (ADF) (1979) and Phillips–Perron (PP) (1988) unit-root tests have been employed for that purpose and the results are summarized in Tables 1. Test for stationarity shows that all variables are integrated of order 1 and thus stationary in difference.

**Table 1: Unit Root Test**

Variables	ADF test statistic (with trend and intercept)		PP test statistic (with trend and intercept)	
	Level	First Difference	Level	First Difference
	e			
<i>ln RGDP</i>	-1.29	-4.14**	-1.34	-3.82**
<i>ln INV</i>	-1.67	-3.91**	-0.85	-3.62**
<i>ln L</i>	-1.91	-3.99**	-1.10	-3.99**
<i>ln OILREV</i>	-1.38	-7.81***	-1.62	-11.20***
<i>SER</i>	-2.18	-5.29***	-1.81	-4.54***

Notes: \*\* and \*\*\* denotesignificancy at 5% 1% levels respectively. The optimal lag structure is determined by SBC

To investigate the presence of long-run relationships among the variables, testing of the bound under Pesaran, et al. (2001) procedure is used. The results of the bound test are given in Table 2. The critical values used in this paper are extracted from Narayan (2004). The calculated F-statistics is 6.19 while upper critical bound at significance level 1% is 5.642. This implies that there is long run relationship among GDP, INV, oil revenues, labor force and education over the period of 1970-2010 in Iran.

**Table 2: Bounds Test Results**

Bound Critical values

F-statistics	Lag	Significance Level	I(0)	I(1)
6.19	2	1%	4.324	5.642
		5%	3.116	4.094
		10%	2.596	3.474

The next stage of the procedure would be to estimate the coefficients of the long-run relations and the associated error correction model (ECM) using the ARDL approach. The optimal lags on variables were selected by the Schwartz Bayesian Criterion (SBC) and turned out to be the ARDL (1, 0, 1, 1, 1). The long-run estimated coefficients are shown in the Table 3. As can be seen, all the coefficients are significant. One percent rise in INV is expected to increase GDP per capita by just 0.17 percent. Although INV appears with the expected positive sign and significant, but the coefficient is small in size. The labor force and the quality of human capital have been the main ingredients for economic growth. The variable of oil revenues has also the expected positive sign.

**Table 3. Estimated long run coefficients based on ARDL approach**

Regressor	Coefficient	p-value
<i>constant</i>	4.20	0.00
<i>ln INV</i>	0.17	0.02
<i>ln L</i>	0.34	0.00
<i>ln OILREV</i>	0.23	0.00
<i>SER</i>	0.32	0.01

The results of error correction model, reported in Table 4. The short-run coefficients are less than the long-run ones. The results suggest that the short-run impact of labor force and human capital on the economic growth are small and insignificant. The coefficients for the other explanatory variables have the expected sign and are significance. Moreover, the coefficient of the ECM is negative and strongly significant at 1% level. This corroborates the existence of a stable long-run relationship and points to a long-run cointegration relationship among variables. The ECM represents the speed of adjustment to restore equilibrium in the dynamic model following a disturbance. The coefficient of the ECM is around -0.71, implying that a deviation from the long-run equilibrium is corrected by 71% after each year.

The diagnostic tests e.g., Lagrange Multiplier (LM) for serial correlation, ARCH effects, normality of residual terms, white heteroskedasticity and Ramsy RESET for functional form reported in Table 5 suggest that the short-run model passes all diagnostic tests. We find no evidence of serial correlation, autoregressive conditional heteroskedasticity and white heteroskedasticity. The residual terms are normally distributed and the functional form of the model appears well specified.

**Table4. Error correction representation for the selected ARDL model**

Regressor	Coefficient	p-value
$\Delta \ln INV$	0.57	0.00
$\Delta \ln L$	0.14	0.19
$\Delta \ln OILREV$	0.19	0.00
$\Delta SER$	0.15	0.26
$ECM$	-0.71	0.00

*Serial Correlation LM = 0.81 (0.72)*

*ARCH Test = 0.56 (0.50)*

*Normality Test = 1.49(0.49)*

*Heteroscedisticity Test = 0.88 (0.72)*

*Ramsey RESET Test = 1.81 (0.39)*

Notes: The probability values for the diagnostic tests are given in parenthesis

The plots of the cumulative sum of recursive residuals (CUSUM) and cumulative sum of squares of recursive residuals (CUSUMSQ) stability tests as shown respectively in figures 1 and 2 indicate that all the coefficients of estimated model are stable over the study period as they fall within the critical bounds.

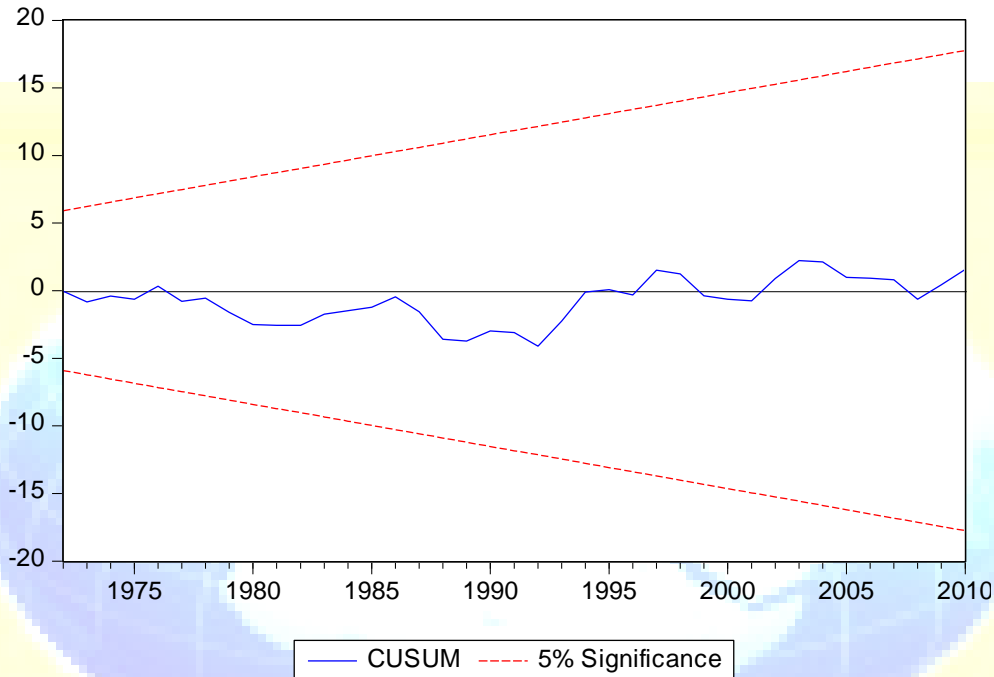


Figure 1: CUSUM Plots for Stability Tests

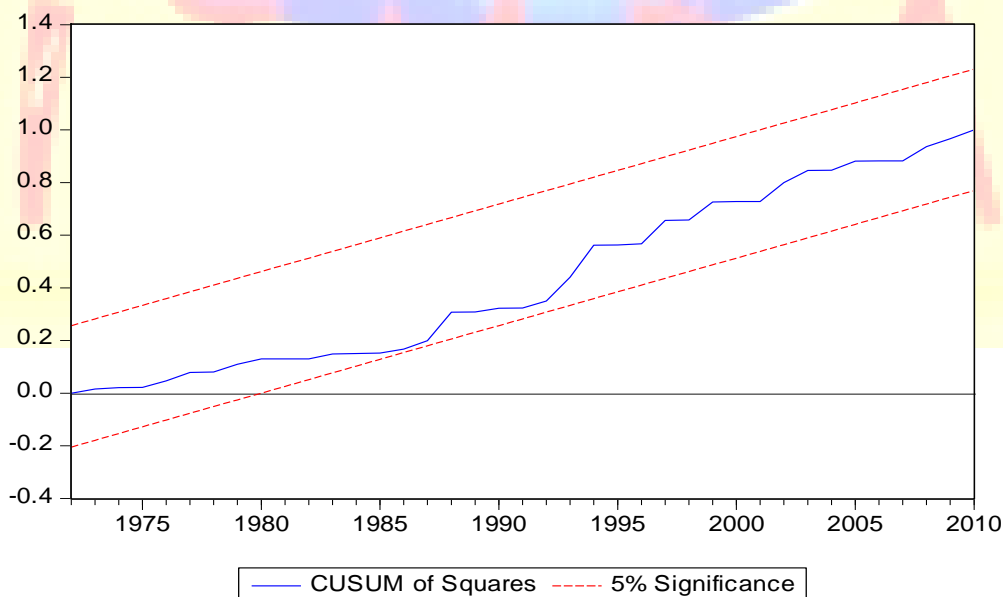


Figure 2: CUSUMSQ Plots for Stability Tests



#### 4. Conclusion

This paper has investigated the determinants of economic growth with an emphasis on the effects of investment in Iran using annual data for the period 1970-2008 applying autoregressive distributed lag (ARDL) approach. According to the results, we found a cointegration relationship among real GDP, investment, labor force, oil revenues and education. Estimating error correction model revealed that the speed of adjustment to restore equilibrium is -0.71 which confirms that there is a stable long-run relationship. Compared to the other variables, labor force and human capital have the most important effect on economic growth in long-run. In short-run, however, the variables of labor force and education do not have significant effects on economic growth, explaining just a small part of economic growth. But the investment and oil revenues appear to play a more important role in short-run growth. Therefore, it does not seem that investment would contribute to economic growth particularly in long-run.

With regard to the important impact of labor force and human capital on economic growth, training skilled labor and professionals in various sectors of the economy and increasing labor productivity can be an essential step in order to stimulate higher long-run growth. In this regard, it is necessary to develop some appropriate policies.

The results show that oil revenue has a more important role in long-run economic growth than investment. To achieve sustainable growth in the future, given the high dependence of Iran economy on oil revenues, it must take policy measures that substantially enlarge and diversify their economic base.

#### Acknowledgements

The authors would like to acknowledge the financial support of university of Tehran for this research under the grant number 4401012/1/19.

**References**

1. Agrawal, P., 2000. Savings, Investment and Growth in South Asia. Indira Gandhi Institute of Development Research.
2. Blomstrom, M., R. E. Lipsey, and M. Zejan, 1996, Is fixed investment the key to economic growth? Quarterly Journal of Economics February, 269-276.
3. De Long, J. B. and L. Summers, 1991, Equipment investment and economic growth. Quarterly Journal of Economics CVI 445-502.
4. De Long, J. B. and L. Summers, 1992, Equipment investment and economic growth: How strong is the nexus? Brookings Papers on Economic Activity, 157-211.
5. Dickey, D., Fuller W., 1979. Distribution of the Estimators for Autoregressive Time Series with a Unit Root, Journal of the American Statistical Association 74, pp. 427-431.
6. Dickey, D. A., W. A. Fuller, 1981. Likelihood Ratio Statistics for Autoregressive Time Series with a Unit Root, Econometrica 49, pp. 1057-1072.
7. Engle, R.F. and Granger, C.W.J. (1987) Cointegration and error-correction: Representation, estimation and testing. Econometrica, 55(2), pp. 987-1008.
8. Granger, C. W. J., 1969. Investigating Causal Relations by Econometric Models and Cross-spectral Methods, Econometrica 37 (3), pp. 424-438.
9. Granger, C. W. J., 1988. Some recent developments in a concept of causality, Journal of Econometrics 39, pp. 199–211.
10. Gregory, Allan W., Hansen, Bruce E., 1996. Residual-based Tests for Cointegration in Models with Regime Shifts, Journal of Econometrics, Elsevier 70(1), pp. 99-126.
11. Hatanaka M., 1996. Time-Series-Based Econometrics: Unit Roots and Cointegration, Oxford University Press.
12. Johansen, S., 1988. Statistical Analysis of Cointegration Vectors, Journal of Economic Dynamics and Control 12 (2–3), pp. 231–254.
13. Johansen, S., 1991. Estimation and Hypothesis Testing of Cointegration Vectors in Gaussian Vector Autoregressive Models, Econometrica 59(6), pp. 1551–1580.
14. Johansen, S., Juselius, K., 1990. Maximum Likelihood Estimation and Inference on Cointegration—with Applications to the Demand for Money, Oxford Bulletin of Economics and Statistics 52 (2), pp. 169–210.
15. Johansen, S., Juselius, K., 1992. Testing structural hypotheses in a multivariate cointegration analysis of the PPP and the UIP for UK, Journal of Econometrics 53, 211–244.

16. Kaldor, N., 1956. Alternative Theories of Distribution, *Review of Economic Studies* 23 (2), pp. 83-100.
17. Kunitomo, N. (1996) Tests OF Unit roots and Cointegration Hypotheses in Econometric Models, 47(1), pp. 79-109.
18. Lewis, W.A. (1955). *The Theory of Economic Growth*, Irwin, Homewood
19. Lutkepohl, H., 2004, *Vector Autoregressive and Vector Error Correction Model*, in Lutkepohl, H. and M. Kratzig (ed.), *Applied Time Series econometrics*, Cambridge University Press.
20. Masih, A. M. M., and R. Masih, 1996. Energy consumption, real income and temporal causality: results from a multi-country study based on cointegration and error-correction modeling techniques, *Energy Economics* 18, pp. 165–183
- Asafu-Adjaye, J., 2000. The relationship between energy consumption, energy prices and economic growth: time series evidence from Asian developing countries, *Energy Economics* 22, pp. 615–625.
21. Pesaran, M. H., & Pesaran B., 1997. *Working with Microfit 4.0: Interactive Econometric Analysis* New York: Oxford University Press.
22. Pesaran, M.H., Shin, Y., & Smith, R.J., 2001. Bounds Testing Approaches to The Analysis of Level Relationships, *Journal of Applied Econometrics* 16, 289–326.
23. Phillips, P.C.B., & Perron, P., 1988. Testing for a Unit Root in Time Series Regression, *Biometrika* 75, 335-46.
24. Pahlavani, M., Wilson, E. J., and A. Valadkhani, 2006. Identifying major structural breaks in the Iranian macroeconomic variables, *International Journal of applied Business and Economic Research* 4(1), pp. 23-44.
25. Perron, p., 1989. The Great Crash, The Oil Price Shock and The Unit Root Hypothesis, *Econometrica* 57, pp. 1361-1401.
26. Phillips, P.C.B., P. Perron, 1988. Testing for a unit root. *Biometrika* 75, pp. 335-346.
27. Samuelson, P., and Modigliani, P., 1966. The Passinetti Paradox in Neo-classical and More General Models, *Review of Economic Studies* 33, pp. 269-301
28. Sinha, D., and Sinha, T., 1998. An Exploration of the Long-run Relationship between Saving and Investment in the Developing Economies: A Tale of Latin American Countries, *Journal of Post Keynesian Economics* 20 (3), pp. 435-443.
29. Tang, T., 2003. Japanese Aggregate Import Demand Function: Reassessment from Bound Testing Approach, *Japan and the World Economy* 15 (4), 419-436.
30. Zivot, E., Andrews, D., 1992. Further Evidence of the Great Crash, the Oil-price Shock and the Unit Root Hypothesis, *Journal of Business and Economic Statistics* 10, pp.251-70.